

# Lecture 11:

Recall: A martingale is a stochastic process  $X_0, X_1, X_2, \dots$  with respect to a filtration  $\tilde{\mathcal{F}}_1, \tilde{\mathcal{F}}_2, \dots$  if  $X_n$  is adapted to  $\tilde{\mathcal{F}}_n$  and  $\mathbb{E}(X_n | \tilde{\mathcal{F}}_{n-1}) = X_{n-1}$ .

$$\left[ \begin{array}{l} \text{Super martingale} \quad \leftrightarrow \quad \leq \\ \text{Sub martingale} \quad \leftrightarrow \quad \geq \end{array} \right]$$

Usually we take  $\tilde{\mathcal{F}}_n = \sigma(X_1, X_2, \dots, X_n)$  but not always!

Remark: let  $m < n$ . For every martingale, we have

$$\begin{aligned} \mathbb{E}(X_n | \tilde{\mathcal{F}}_m) &= \mathbb{E}(\mathbb{E}(\dots \mathbb{E}(\underbrace{\mathbb{E}(X_n | \tilde{\mathcal{F}}_{n-1})}_{= X_{n-1}} | \tilde{\mathcal{F}}_{n-2}) \dots | \tilde{\mathcal{F}}_m) | \tilde{\mathcal{F}}_m) \\ &= \mathbb{E}(\mathbb{E}(\dots \underbrace{\mathbb{E}(X_{n-1} | \tilde{\mathcal{F}}_{n-2})}_{X_{n-2}} \dots | \tilde{\mathcal{F}}_m) | \tilde{\mathcal{F}}_m) \\ &\quad \vdots \\ &= \mathbb{E}(X_{m+1} | \tilde{\mathcal{F}}_m) = X_m. \end{aligned}$$

Definition A **pre-visible process** is a sequence  $C_1, C_2, \dots$  of random variables, such that  $C_n$  is  $\mathcal{F}_{n-1}$ -measurable for all  $n$ .

Let  $C_n$  be a pre-visible process. The **martingale transform** of  $X$  by  $C$  is

$$(C \cdot X)_n = \sum_{k=1}^n C_k (X_k - X_{k-1}).$$

In particular, if  $C_k = 1$  for all  $k$ , then  $(C \cdot X)_n = X_n - X_0$ .

Proposition: If  $C$  is a bounded pre-visible process with  $|C_n(\omega)| \leq K$  for all  $n$  and  $\omega \in \Omega$ , then  $(C \cdot X)_n$  is a martingale if  $X_n$  is. If  $C$  is also non-negative, then  $(C \cdot X)_n$  is a sub-/super martingale whenever  $X_n$  is.

Proof: We have

$$\begin{aligned} & E((C \cdot X)_n - (C \cdot X)_{n-1} \mid \mathcal{F}_{n-1}) \\ &= E(\underbrace{C_n}_{\dots\dots\dots} (X_n - X_{n-1}) \mid \mathcal{F}_{n-1}) \\ &= C_n (E(X_n \mid \mathcal{F}_{n-1}) - E(X_{n-1} \mid \mathcal{F}_{n-1})) \\ &= C_n (E(X_n \mid \mathcal{F}_{n-1}) - X_{n-1}) \end{aligned}$$

$$\begin{cases} = 0 & \text{if } X_n \text{ is a martingale} \\ \geq 0 & \text{if } C_n \geq 0 \text{ and } X_n \text{ is a submartingale} \\ \leq 0 & \text{if } C_n \geq 0 \text{ and } X_n \text{ is a supermartingale} \end{cases}$$

□

### Stopping times

A **stopping time** is a random variable  $T$  with values in  $\{0, 1, 2, \dots, \infty\}$  and the property that  $\{T \leq n\} = \{\omega \in \Omega : T(\omega) \leq n\} \in \mathcal{F}_n$  for all  $n$ . Equivalently,  $\{T = n\} \in \mathcal{F}_n \forall n$ . This follows from  $\{T \leq n\} = \{T \leq n-1\} \cup \{T = n\}$

Examples: • All constants are stopping times

- "First occurrence": for example:

$$T = \min \{ n : X_n = 0 \} \text{ for an adapted process } X_n$$

- If  $S, T$  are stopping times, then so are

$$\min \{ S, T \} = S \wedge T \quad \text{"either stopped"}$$

$$\text{and } \max \{ S, T \} = S \vee T \quad \text{"both stopped"}$$

- "Counting": for example, set

$$N_n = \text{number of indices } k \leq n \text{ with } X_k = 0$$

$$T = \min \{ n : N_n = 10 \}$$

Example: The following are (generally)

not stopping times

$$T = \max \{ n : N_n = 0 \} \quad (\text{we cannot determine whether } N_k = 0 \text{ for } k > n)$$

Also

$$T = \min \left\{ n : X_n = \sup_k X_k \right\} \quad \left( \sup_k X_k \text{ not measurable wr.t. } \mathcal{F}_n. \text{ Could be larger later} \right)$$

## Stopped processes:

Let  $X_n$  be an adapted process and  $T$  a stopping time with respect to a given filtration. The stopped process  $X^T$  is

$$X_n^T(\omega) = X_{n \wedge T(\omega)}(\omega) = \begin{cases} X_{T(\omega)}(\omega) & \text{if } n \geq T(\omega) \\ X_n(\omega) & \text{if } n < T(\omega) \end{cases}$$

Theorem: If  $X_n$  is a martingale / supermartingale / submartingale, then so is  $X_n^T$ . In particular, for every  $n$ ,  $E(X_{T \wedge n}) \stackrel{\leq}{=} E(X_0)$  <sup>super-</sup> <sub>sub-</sub> martingale.

Proof: Note that  $C_n^T = \underbrace{I_{\{n \leq T\}}}_{\text{"not yet stopped at time } n-1"} = 1 - I_{\{T \leq n-1\}}$  is pre-visible.

$$\text{We have } (C^T \cdot X)_n = \sum_{k=1}^n C_k^T (X_k - X_{k-1})$$

$$= \sum_{k=1}^n I_{\{k \leq T\}} (X_k - X_{k-1}) = \sum_{k=1}^{T \wedge n} (X_k - X_{k-1})$$

$$= X_{T \wedge n} - X_0. \quad \text{So } X_{T \wedge n} \text{ is a martingale.}$$

Since  $E(E(X|\mathcal{F}^T)) = E(X)$ , the second conclusion follows  $\square$

So for every fixed  $n$ ,  $E(X_{T \wedge n}) = E(X_0)$ .

Is it true that  $E(X^T) = E(X_0)$ ?

In general, NO!

Example: Consider the martingale

$$X_0 = 1, \quad X_n = \begin{cases} 2X_{n-1} & \text{prob. } \frac{1}{2} \\ 0 & \text{prob. } \frac{1}{2} \end{cases}$$

Let  $T = \min\{n : X_n = 0\}$ . Clearly,  $E(X^T) = 0 \neq E(X_0)$ .

Example: Consider the simple random walk

$$X_0 = 0, \quad X_n = \begin{cases} X_{n-1} + 1 & \text{prob. } \frac{1}{2} \\ X_{n-1} - 1 & \text{prob. } \frac{1}{2} \end{cases}$$

Define  $T = \min\{n : X_n = 1\}$ . This is a stopping time and one can show that  $T < \infty$  a.s.

Hence  $E(X^T) = 1 \neq E(X_0)$ .

However, under simple conditions  $E(X^T) = E(X_0)$

## Doob's Optional Stopping Theorem

Let  $T$  be a stopping time, and let  $X$  be a <sup>super-</sup>sub-  
martingale. Suppose one of the following hold:

- i)  $T$  is bounded (almost surely)
- ii)  $X_n$  is bounded and  $T < \infty$  for a.e.  $\omega \in \Omega$ .
- iii)  $E(T) < \infty$  and  $|X_n(\omega) - X_{n-1}(\omega)| \leq K$  for some fixed  $K$ , all  $n$  and almost every  $\omega \in \Omega$ .

Then,  $E(X^T) \begin{cases} \leq \\ = \\ \geq \end{cases} E(X_0)$  super-  
martingale.  
sub-

Proof: (i) If  $T$  is bounded by  $N$  (a.s.)

we have  $T \wedge N = T$  (a.s.) and so

$$E(X^T) = E(X_{T \wedge N}) \begin{cases} \leq \\ = \\ \geq \end{cases} E(X_0).$$

(ii) We have  $E(X_{T \wedge n}) \begin{cases} \leq \\ = \\ \geq \end{cases} E(X_0)$  for fixed  $n$ .

Since  $X$  is bounded we can use dominated

convergence:

$$\mathbb{E}(X_0) = \lim_{n \rightarrow \infty} \mathbb{E}(X_{T_{1n}}) = \mathbb{E}\left(\lim_{n \rightarrow \infty} X_{T_{1n}}\right) - \mathbb{E}(X_T).$$

iii) We have

$$\begin{aligned} |X_{T_{1n}} - X_0| &= \left| \sum_{k=1}^{T_{1n}} (X_k - X_{k-1}) \right| \\ &\leq \sum_{k=1}^{T_{1n}} |X_k - X_{k-1}| \leq K \cdot (T_{1n}) \leq KT. \end{aligned}$$

So  $\mathbb{E}(KT) = K \mathbb{E}(T) < \infty$  and we can apply DCT as above.  $\square$

Corollary: If  $X_n$  is a nonnegative supermartingale and  $T$  an (almost surely) finite stopping time, then

$$\mathbb{E}(X^T) \leq \mathbb{E}(X_0)$$

Proof: By Fatou's lemma,

$$\begin{aligned} \mathbb{E}(X^T) &= \mathbb{E}\left(\lim_{n \rightarrow \infty} X_{T_{1n}}\right) \quad (\text{why does it exist?}) \\ &= \mathbb{E}\left(\liminf_{n \rightarrow \infty} X_{T_{1n}}\right) \\ &\leq \liminf_{n \rightarrow \infty} \mathbb{E}(X_{T_{1n}}) \leq \mathbb{E}(X_0) \end{aligned}$$