

Strong law of large numbers (First version)

Theorem: Let X_1, X_2, \dots be a sequence of independent random variables. Suppose that $\mathbb{E}(X_i) = 0$ and $\mathbb{E}(X_i^4) \leq K$ for all i and some uniform constant $K > 0$. Then,

$$\frac{X_1 + X_2 + \dots + X_n}{n} \rightarrow 0 \text{ almost surely.}$$

Proof: We consider the fourth moment.

$$\begin{aligned} S_n &= X_1 + X_2 + \dots + X_n \\ \mathbb{E}(S_n^4) &= \mathbb{E}(X_1 + \dots + X_n)^4 \\ &= \mathbb{E}(X_1^4) + \dots + \mathbb{E}(X_n^4) \\ &\quad + 4 (\mathbb{E}(X_1 X_2^3) + \mathbb{E}(X_1^3 X_2) + \dots) \\ &\quad + 6 (\mathbb{E}(X_1^2 X_2^2) + \dots) \\ &\quad + 12 (\mathbb{E}(X_1^2 X_2 X_3) + \dots) \\ &\quad + 24 (\mathbb{E}(X_1 X_2 X_3 X_4) + \dots) \end{aligned}$$

We can use independence for mixed terms

$$\text{such as } \mathbb{E}(X_1^2 X_2 X_3) = \mathbb{E}(X_1^2) \mathbb{E}(X_2) \mathbb{E}(X_3) = 0.$$

We are left with terms of form

1) $E(X_i^4)$ but they are bounded by K .

2) $E(X_i^2 X_j^2) \leq \sqrt{E(X_i^4) E(X_j^4)} \leq K$ for all i, j .

by Cauchy-Schwarz,

So, $E(S_n^4) \leq nK + G \cdot \binom{n}{2} K$

$$= K(n + 3n(n-1)) = K_n(3n-2) \leq 3K_n^2$$

So $E\left(\left(\frac{S_n}{n}\right)^4\right) \leq \frac{3K}{n^2}$

$$\Rightarrow E\left(\sum_{n=1}^{\infty} \left(\frac{S_n}{n}\right)^4\right) \leq 3K \sum_{n=1}^{\infty} \frac{1}{n^2} < \infty.$$

So with probability 1, $\sum_{n=1}^{\infty} \left(\frac{S_n}{n}\right)^4$ converges

and with probability 1 $\frac{S_n}{n} \rightarrow 0$. \square

Special case: If X_1, X_2, \dots are independent and identically distributed with $E(X_i^4) < \infty$,

then $Y_i = X_i - E(X_i)$ has expectation 0.

Y_1, Y_2, \dots satisfy the conditions above and

$$\frac{Y_1 + \dots + Y_n}{n} \rightarrow 0, \frac{(X_1 - m) + \dots + (X_n - m)}{n} \rightarrow 0$$

and $\frac{X_1 + X_2 + \dots + X_n}{n} \rightarrow m$, when $m = E(X_i)$.

We can derive the distance to $\mu = E(X)$ by means of Chebyshov's inequality.

$$P(|X - \mu| \geq c) \leq \frac{\text{Var}(X)}{c^2}.$$

If is a special case of Markov's inequality:

$$P(|X - \mu|^2 \geq c) \leq \frac{1}{c^2} E(|X - \mu|^2) = \frac{\text{Var}(X)}{c^2}.$$

Applying this to $S_n = X_1 + X_2 + \dots + X_n$, where X_1, X_2, \dots are independent identically distributed (iid) random variables with $\mu = E(X)$ and $\sigma^2 = \text{Var}(X_i) < \infty$ gives

$$E(S_n) = E(X_1) + \dots + E(X_n) = n\mu,$$

$$\text{Var}(S_n) = \text{Var}(X_1) + \dots + \text{Var}(X_n) = n\sigma^2 \text{ by independence}$$

$$\text{Further } \text{Var}\left(\frac{S_n}{n}\right) = \text{Var}\left(\frac{S_n - n\mu}{n}\right) = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$$

$$\text{So } P\left(\left|\frac{1}{n}S_n - \mu\right| \geq c\right) \leq \frac{\sigma^2}{n \cdot c} \rightarrow 0 \text{ as } n \rightarrow \infty$$

Since c was arbitrary we have

$$\frac{S_n}{n} \rightarrow \mu \quad \text{in probability.}$$

Conditional Expectation

Simple Example: Throw a die. All outcomes $1, 2, \dots, 6$ are equally likely. Write X for the r.v. that gives the outcome.

We have $P(X \leq 3) = \frac{3}{6} = \frac{1}{2}$.

Suppose we additionally know the outcome is even or odd.

The conditional probabilities of the event

$\{X \leq 3\}$ are

$$P(X \leq 3 | X \text{ odd}) = \frac{P(\{X \leq 3\} \cap \{X \text{ odd}\})}{P(\{X \text{ odd}\})} = \frac{2}{3}$$

$$P(X \leq 3 | X \text{ even}) = \frac{P(\{X \leq 3\} \cap \{X \text{ even}\})}{P(\{X \text{ even}\})} = \frac{1}{3}$$

We can also obtain conditional expectations:

$$E(X | X \text{ even}) = \frac{2+4+6}{3} = 4$$

$$E(X | X \text{ odd}) = \frac{1+3+5}{3} = 3$$

Generally, we define conditional expectations with respect to σ -algebras

Definition (Theorem): Let (Ω, \mathcal{F}, P) be a probability space and $\mathcal{G} \subseteq \mathcal{F}$ a sub σ -algebra. Let X be an integrable random variable.

Then exists a random variable $Y = Y(\omega)$ with the following properties:

(1) Y is \mathcal{G} -measurable

(2) Y is integrable

(3) For all $G \in \mathcal{G}$: $\int_G Y dP = \int_G X dP$

Moreover Y is almost surely unique.

For any Y, Y' satisfying (1)-(3), $P(Y=Y')=1$.

This random variable is called the conditional expectation of X wrt. \mathcal{G} and we write

$$Y(\omega) = \underline{E}(X | \mathcal{G}) (= E(X | \mathcal{G})(\omega)).$$

If \mathcal{G} is generated by random variables we write

$$\mathbb{E}(X | \mathcal{Z}) \text{ instead of } \mathbb{E}(X | \sigma(\mathcal{Z}))$$

$$\mathbb{E}(X | Z_1, Z_2, \dots, Z_n) \text{ --- } \mathbb{E}(X | \sigma(Z_1, \dots, Z_n)).$$

Example: In our die throwing example

$\tilde{\rho} = P(\{1, \dots, 6\})$, $\mathcal{G} = \{\emptyset, \{1, 3, 5\}, \{2, 4, 6\}, \Omega\}$

\mathcal{G} measurability $\Rightarrow Y$ must be constant on $\{1, 3, 5\}$ and $\{2, 4, 6\}$, respectively.

$$\text{Say } Y(\omega) = \begin{cases} a & \text{for } \omega \in \{1, 3, 5\} \\ b & \text{for } \omega \in \{2, 4, 6\} \end{cases}.$$

$$1) \int_{\emptyset} Y dP = \int_{\emptyset} X dP \quad \text{Trivial } 0=0.$$

$$2) \int_{\{1, 3, 5\}} Y dP = \int_{\{1, 3, 5\}} X dP$$

$$\Rightarrow a P(\{1, 3, 5\}) = 1 P(\{1\}) + 3 P(\{3\}) + 5 P(\{5\})$$

$$\Rightarrow \frac{1}{2} a = \frac{1+3+5}{6} \Rightarrow a = 3$$

$$(3) \int_{\{2,4,6\}} Y dP = \int_{\{2,4,6\}} X dP$$

$$\Rightarrow \frac{1}{2} b = \frac{2+4+6}{6} \Rightarrow b=4.$$

$$(4) \int_{\Omega} Y dP = \int_{\Omega} X dP$$

Sum of (2) & (3) so also satisfied.

Remark: The "ordinary" expectation is the special case $\mathcal{G} = \{\emptyset, \Omega\}$. Then Y is constant over Ω and

$$\int_{\emptyset} X dP = \int_{\emptyset} Y dP \quad (\text{trivial})$$

$$\int_{\Omega} X dP = \int_{\Omega} Y dP \Rightarrow Y(\omega) = \int_{\Omega} X dP = E(X).$$

We may interpret or algebras as "knowledge" of an event and will investigate

sequences X_1, X_2, \dots of integrable random variables:

Central theme:

Defⁿ Let X_1, X_2, \dots be integrable random variables. The sequence is a martingale

if $E(X_{n+1} | \sigma(X_1, X_2, \dots, X_n)) = X_n$.

Informally, the expectation of the $(n+1)$ -th random variable conditioned on "knowing" outcomes X_1, \dots, X_n is equal to the last observed value.

Properties of conditional expectation:

Th^m We have

$$(1) E(E(X | \mathcal{G})) = E(X)$$

$$(2) \text{ If } X \text{ is } \mathcal{G} \text{ measurable, then } E(X | \mathcal{G}) = X \text{ a.s.}$$

$$(3) \text{ Linearity: } E(aX + bY | \mathcal{G}) = aE(X | \mathcal{G}) + bE(Y | \mathcal{G}) \text{ a.s.}$$

$$(4) \text{ Positivity: If } X \geq 0 \text{ a.s. then } E(X | \mathcal{G}) \geq 0 \text{ a.s.}$$

Proof :

(1) Since $\Omega \in \mathcal{G}$, we have

$$E(E(X|\mathcal{G})) = \int_{\Omega} E(X|\mathcal{G}) dP = \int_{\Omega} X dP = E(X)$$

(2) X satisfies all conditions in definition / theorem

(3) Note that $aE(X|\mathcal{G}) + bE(Y|\mathcal{G})$

is \mathcal{G} -measurable. The condition then follows by linearity of integration.

(4) Suppose $Y = E(X|\mathcal{G})$ is negative with positive probability : $P(Y < 0) > 0$.

Then $\exists n$ s.t. $P(Y \leq -\frac{1}{n}) > 0$. This is hence a \mathcal{G} measurable set and

$$\begin{aligned} \underbrace{\int_{\{Y \leq -\frac{1}{n}\}} Y dP}_{\leq -\frac{1}{n} P(\{Y \leq -\frac{1}{n}\}) < 0} &= \int_{\{Y \leq -\frac{1}{n}\}} X dP \\ &\geq 0 \end{aligned}$$

A contradiction.

□

Results on convergence carry over: non-neg.

Th^m (1) If X_1, X_2, \dots is a sequence of random variables such that $X_n \uparrow X$, then we also have

$$E(X_n | \mathcal{G}) \uparrow E(X | \mathcal{G}). \quad (\text{MCT})$$

(2) If X_1, X_2, \dots is a sequence of random variables s.t. $|X_n| \leq Y$ for some integrable Y , and $X_n \rightarrow X$, then also $E(X_n | \mathcal{G}) \rightarrow E(X | \mathcal{G})$ (DCT)

(3) If X_1, X_2, \dots is any sequence of non-negative random variables, then

$$E(\liminf_{n \rightarrow \infty} X_n | \mathcal{G}) \leq \liminf_{n \rightarrow \infty} E(X_n | \mathcal{G}). \quad (\text{Fatou})$$